

MODELS, METHODS & APPLICATIONS OF ECONOMETRICS
EA PCB Phillips Sackwell 1993
17B139 M626

DFA:
1986-87

60 DISTRICTS
3.3 million population
19 TWh sales

23

Economies of Scale in the New Zealand Electricity Distribution Industry

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per capita
annual consumption: 6 MWh

Output: 14 MWh

Introduction

The New Zealand electricity industry is currently undergoing a process of deregulation, and there is considerable interest in its future structure. This industry is divided into three horizontal layers, dealing with the generation, high-voltage transmission, and final distribution of electricity. Although a few distributors generate a significant proportion of their requirements, there has been no trend towards vertical integration. Generation and transmission are currently under the control of the Electricity Corporation of New Zealand (ECNZ), from whom the various regional Electricity Supply Authorities (ESAs) purchase bulk electricity, and transmit it from ECNZ's substations to the end users.¹ This study focuses on these ESAs and examines whether there are economies of scale in the distribution of electricity in New Zealand.

At the time of our sample there were 60 ESAs in New Zealand,² a mountainous country comprising two principal islands, with a population of approximately 3.3 million people. Some 70 percent of the population live in the North Island (44,281 square miles), the Auckland metropolitan area accounting for 840,000 people, while in the South Island (58,093 square miles) the largest urban centre is Christchurch (population 300,000). Constitutionally, the ESAs are of two types - Electric Power Boards (EPBs), which are independent statutory bodies run by boards elected from the area over which the ESA has a franchise; and the Municipal Electricity Departments (MEDs), which are the trading arms of territorial local government and are managed by committees of the relevant local council. Each ESA has an area franchise and an associated obligation to supply electricity without discrimination. Only licensed firms may distribute electricity. Strictly, these franchises do not exclude the operation of competitors, but in practice only one franchise has been issued for each area.

Debate over the scale of electricity distributors in New Zealand has raged for years. The geographical features of the country are unusually relevant, but given the number of consumers and the relatively small annual output (19,444 GWh in 1986/87), it is not surprising that the appropriateness of 60 or so ESAs has been questioned. A Royal Commission headed by Stanton (1959) concluded that the (then 83) ESAs were technically efficient, but would be improved by amalgamations. So 26 supply areas were recommended, but without explicit reference to (economic) scale economies.

The latter were explored to some extent by Jones (1987) and with respect to the Christchurch region by McCutcheon et al. (1987). None of these studies employs any formal econometric analysis.

Several technical and organizational factors influence economies of scale in electricity distribution. Organizational economies may arise at the firm level as a result of staff specialization and staff control costs. Below some size a firm may not be able to employ the optimal resources, while beyond some other size the firm faces increasing costs in controlling these resources. Organizational economies can arise from increasing the specialization of managerial staff - the benefits of a larger firm depend on the gains from having this expertise "in house" in terms of cost and firm-specific knowledge gained. Conversely, organizational diseconomies may develop beyond a certain firm size, perhaps because of increasing communication problems, the difficulty of maintaining consistent objectives, and the potential for managerial "slack."

The principal technical economies are in distribution equipment, which lead to economies of density, economies in capacity expansion, and economies in the provision of capacity to meet peak requirements. An increase in equipment capacity leads to a less than proportionate rise in equipment cost. Larger capacity equipment also yields lower system costs as higher voltage operation lowers system energy losses. These factors contribute to economies of density. As the number of customers, and energy demand, rises for a given area, average cost falls. Supply security can also be improved when density rises. Several low-voltage networks can be interconnected with open switches to provide different flow paths, so that a fault that might otherwise cut off supply can be bypassed. Such benefits are exhausted at a certain scale of operation by the requirement to keep separate the networks supplied by each ECNZ substation. All of this points to the potential for scale economies in this industry - the extent to which such economies are in fact present or are exhausted is, of course, an empirical issue.

Econometric studies of scale economies in electricity industry in other countries focus primarily on generation rather than distribution, and reflect the vertically integrated nature of this industry elsewhere, see for example Christensen and Greene (1976), Betancourt and Edwards (1987), and Sing (1987). Other relevant studies include those of Neuberger (1977), Heuttner and Landort (1978), Avazian et al. (1987), and some of the material discussed by Weiss (1975). A typical finding is that the average cost curves exhibit extensive "flat" regions - i.e. there is a wide range of outputs consistent with essentially constant returns to scale.

In this paper we use a Translog cost model and cross-section data for the 1986/87 financial year to estimate economies of scale in the distribution of electricity in New Zealand. The ESAs have a statutory obligation to supply electricity and are price takers in the purchase of bulk electricity from ECNZ (this being their major operating cost). An econometric model of cost, rather than production, is appropriate given that the firms are cost minimizers rather than profit maximizers.³ The model and data used are described in the next two sections. The fourth section discusses the results, and our conclusions are given in the final section.

The Model

Costs for the firms in this industry are of the form

$$C = f(Y, P, I), \quad (1)$$

where C denotes total cost, Y is output, P is a vector of n input prices, and I is a vector of m additional industry-specific variables. In formulating such a relationship, it is assumed that output and input prices are exogenous, and that (for a given technology) firms adjust input levels so as to minimize costs of production. Given the comments at the end of the last section, and our use of cross-section data, such assumptions seem reasonable in this case. In common with other related studies, we use the Translog function (e.g. Christensen et al. (1971, 1973)) to formalize (1):

$$\begin{aligned} \ln C = & \alpha_0 + \alpha_Y \ln Y + \frac{1}{2} \gamma_{YY} (\ln Y)^2 + \sum_{i=1}^n \gamma_{Yi} \ln Y \ln P_i \\ & + \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n \gamma_{ij} \ln P_i \ln P_j + \sum_{i=1}^n \alpha_i \ln P_i + \sum_{k=1}^m \beta_k \ln I_k \\ & + \frac{1}{2} \sum_{k=1}^m \sum_{\theta=1}^m \beta_{\theta k} \ln I_k \ln I_\theta + \sum_{k=1}^m \beta_{ky} \ln I_k \ln Y \\ & + \sum_{k=1}^m \sum_{i=1}^n \theta_{ki} \ln I_k \ln P_i, \end{aligned} \quad (2)$$

where $\gamma_{ij} = \gamma_{ji}$ and $\beta_{ky} = \beta_{yk}$. The attractions of this functional form are that it is flexible enough to represent quite general production structures; it imposes no restrictions on factor substitutability; and it allows economies of scale to vary with output. To ensure that (2) is consistent with a well-behaved production function, it must be homogeneous of degree one in input prices, which implies

$$\begin{aligned} \sum_{j=1}^n \gamma_{ij} &= \sum_{i=1}^n \sum_{j=1}^n \gamma_{ij} = \sum_{i=1}^n \gamma_{ij} = 0, \\ \sum_{i=1}^n \gamma_{Yi} &= 0, \quad \sum_{i=1}^n \alpha_i = 1, \quad \sum_{i=1}^n \theta_{ki} = 0; \quad k = 1, \dots, m. \end{aligned} \quad (3)$$

The symmetry of the γ_{ij} terms follows from Young's theorem, given that the cost function is continuously twice differentiable. If X_i is the quantity of input i , applying Shephard's lemma, the associated input cost share equations are

$$S_i = (P_i X_i / C) = \alpha_i + \gamma_{Yi} \ln Y + \sum_{j=1}^n \gamma_{ij} \ln P_j + \sum_{k=1}^m \theta_{ki} \ln I_k, \quad i = 1, \dots, n. \quad (4)$$

Adding normally distributed error terms, our full model comprises equation (2) and the equations (4). The model is estimated by joint maximum likelihood (ML), with all of the mentioned parameter restrictions imposed. The well-known singularity of the contemporaneous error covariance matrix associated with such allocation models is allowed for in the usual way - one of the share equations is dropped from the system, the coefficient estimates being invariant to the choice of this equation. Our primary interest is the measurement of scale economies. As is conventional, such economies are defined in terms of the relationship between cost and output along the expansion path. That is, with fixed input prices (and industry-specific variables)

and costs minimized at each output level. Specifically, we define them as the ratio of average to marginal costs:

$$\text{SCALE} \equiv (\partial \ln C / \partial \ln Y)^{-1} = \left[\alpha_Y + \gamma_{YY} \ln Y + \sum_{i=1}^n \gamma_{Yi} \ln P_i + \sum_{k=1}^m \beta_{ky} \ln I_k \right]^{-1}, \quad (5)$$

so there are scale economies (diseconomies) if SCALE is greater (less) than unity. The requirement that SCALE declines monotonically, as Y increases, holds if and only if $\gamma_{YY} > 0$. Given estimates of the parameters, and fixing the values of input prices and the industry-specific variables, setting (5) to unity and solving for Y allows the calculation of the output level at which such a firm's average cost is minimized. Dividing this output level into the total value of industry output suggests the number of such firms that are consistent with average cost minimization.

Data

Our sample comprises 60 cross-section data points for the 1986/87 financial year, the latest available. Output is defined as total retail sales of electricity (kWh). The sample is characterized by a conglomeration of ESAs with small output - 52 of the 60 firms had an annual output of 500 GWh or less. Of the remainder, one (Auckland) had an output nearly seven times as great as this, more than ten times the average output, and only slightly less than the combined outputs of the next three largest ESAs.⁴

Total cost is the simple sum of four input costs; those associated with labor, capital, electricity purchased, and other. The last of these is measured as total reported expenditure in the year in question, less expenditures on labor, capital, and on purchasing and (if applicable) generating electricity, less expenditures relating to the retailing of appliances. Accordingly, these "other" costs essentially relate to maintenance and operation, administration, loan interest, and depreciation. In the case of labor, total salaries and wages were adjusted to exclude amounts associated with any electricity generation or appliance sales, or capitalized in particular capital projects. Figures for the historic value of each ESA's assets were inflated and depreciated to allow for a life of 30 years for most distribution equipment. Different capital items were treated separately with respect to depreciation rates and method of depreciation (diminishing value or straight line). This breakdown distinguished between distribution equipment; distribution and transmission buildings; public lighting; land; offices, stores, and workshops; loose tools, plant, and furniture; motor vehicles; and other capital items. Aggregate capital figures were then constructed. The cost of purchasing electricity includes that associated with purchases from ECNZ, plus any from firms involved in cogeneration or from other ESAs. Generation costs (in the few cases, where relevant) were excluded from the analysis to ensure comparability across ESAs. A commensurate value was calculated and included in total electricity costs in such cases. Such values were calculated taking account of the structure of the bulk supply tariff, which distinguishes between winter zone and anytime demands; between day and night energy rates; and between location in the North and South Islands.

Input prices were defined as follows. The prices of labor and "other" inputs are

the associated expenditures divided by the number of employees. Three possible definitions of the price of capital inputs were considered: total capital value divided by circuit kilometers of distribution line; the value of capital divided by the combined amount of circuit kilometers of distribution line plus kVA transformer capacity, expressed as circuit kilometer "equivalents"; a constant capital price across ESAs.⁵ Two alternative measures of the price of electricity purchased were considered: total electricity cost divided by total electricity purchased and generated (cents/kWh); and constant values of 2.8084 cents/kWh for South Island ESAs, and 3.4318 cents/kWh for North Island ESAs, reflecting the average energy components in the wholesale electricity tariff.

Finally, we consider up to five industry-specific variables. Two of these are dummy variables for the "type" of ESA ($I_1 = 1$, EPB; $= 0$, MED); and for ESA location ($I_2 = 1$, North Island; $= 0$, South Island). We also considered the load factor (I_3) and density. Two possible measures of the latter were considered - $I_4 =$ number of consumers per square kilometer of licensed area; or $I_4 =$ number of consumers per circuit kilometer of distribution line. Also considered was a regional dummy variable ($I_5 = 1$, urban; $= 0$, rural). Further details of the data and their construction are given by Wyatt et al. (1989).

Results

Joint ML estimation and the other computations were undertaken with the TSP package (Hall et al. 1988). Equation (2) was constrained to avoid the multicollinearity that would otherwise arise because $\ln(I_i) = (\ln(I_i))^2$ ($i = 1, 2, 5$). Not all of $\beta_1, \beta_2, \beta_3, \beta_{11}, \beta_{22}$, and β_{33} are identifiable in our model and so the relevant terms were coded as $1.5\beta_i \ln(I_i)$ ($i = 1, 2, 5$). When all of the industry-specific variables (however defined) were included in the model, the results were economically implausible: the estimates of γ_{yy} were negative. A sequence of nested model tests revealed the load factor to be the main source of this problem,⁶ although when it is removed from the model we were still unable to obtain plausible results with a significant rural-urban effect. Accordingly, we discarded these two regressors. A formal model-selection procedure was then used to determine the final specification of the model. Akaike's information criterion (AIC) was used to discriminate between the non-nested specifications that arise with the alternative definitions of density and the prices of capital and electricity purchased. With the variable definitions fixed, a sequence of nested likelihood ratio tests (LRTs) was used to determine whether the model should be simplified further by deleting other industry-specific variables. These asymptotic tests were applied in the manner described by Mizon (1977) so as to control their true size.

This procedure favored the first of the measures of density and electricity price, and the second measure of capital price. With these definitions, the results of our LRTs appear in table 23.1. Two non-nested specifications cannot be rejected, so the final selection is again based on AIC, resulting in the retention of density and the "type" of ESA as the industry-specific variables. The lack of significance of the North-South dummy apparently reflects the fact that the electricity purchase price variable is already capturing the relevant effects. The ML estimates of the model's parameters appear in table 23.2.

Table 23.1 Tests of nested hypotheses

Model specification tests		LRT*	ν	AIC
<i>Industry-specific variables included</i>				
<i>Maintained hypothesis</i>				
<i>Restricted hypothesis</i>				
(I_1, I_2, I_4)	(I_1, I_2)	N.A. ^b	7	-
(I_1, I_2, I_4)	(I_2, I_4)	11.97	-	-16.97
(I_2, I_4)	I_2	N.A. ^b	6	-
(I_2, I_4)	I_4	20.26 ^c	6	-
(I_1, I_2, I_4)	(I_1, I_4)	8.76	7	-17.03
(I_1, I_2, I_4)	I_1	23.46 ^c	6	-
(I_1, I_4)	I_1	N.A. ^b	-	-
<i>Homotheticity test^d</i>				
<i>Maintained hypothesis</i>				
<i>Restrictions</i>				
$\gamma_{y1} = \gamma_{y2} = \gamma_{y3} = 0$				
See table 23.2		LRT	ν	
		20.40 ^c	3	

* LRT is asymptotically χ^2 with ν degrees of freedom.
 b Not acceptable. Estimated coefficients conflict with prior economic theory.
 c Significant at the 1 percent level or higher.
 d Homogeneity restrictions are nested within homotheticity restrictions, so the former are rejected. Testing separately for homogeneity also leads to rejection: LRT = 21.86 ($\nu = 6$).

Each variable enters the model in several forms. The overall significance of the industry-specific variables has been established via the LRTs. Individually, many of the parameters are also significant. Signing the parameters is not trivial. The relevant issue is the sign of the partial derivative of C with respect to the variable of interest. Equivalently, we can sign the associated elasticity, which is more easily derived, given the form of (2). For example, we expect $\partial \ln C / \partial \ln Y (= 1/SCALE)$ and $\partial \ln C / \partial \ln P_i (= S_i)$ ($i = 1, \dots, 4$) to be positive; and in the case of "density", $\partial \ln C / \partial \ln I_4 < 0$. The anticipated sign for the "ESA-type" dummy variable is ambiguous. Using the estimated parameters from table 23.2, we have evaluated each of these derivatives at each point in the sample. There are no exceptions to the anticipated signs, and the estimated shares are all positive fractions. Given the imposition of the restriction $\sum_{i=1}^4 \gamma_{yi} = 0$, homotheticity of the underlying production function would imply $\gamma_{y1} = \gamma_{y2} = \gamma_{y3} = 0$, homotheticity of the additional restrictions $\gamma_{yy} = \beta_{yy} = \beta_{4y} = 0$ implies homogeneity of the production function. Homotheticity and homogeneity are both rejected by LRTs (see table 23.1), and so are not imposed in the following analysis. Ordering the data by increasing value of output, there is no evidence of heteroskedasticity in the residuals. However, to be conservative, all reported standard errors and tests are based on White's (1980) heteroskedasticity-consistent covariance matrix estimator.

Estimated scale economy measures appear in table 23.3. These are obtained from (5) in two ways - first, as is conventional, SCALE is calculated with all variables except output fixed (here, to their sample means); and secondly, SCALE is calculated at each individual sample point.⁷ The latter figures are interesting, but the former are of primary importance and only on these can cross-firm comparisons be based. The point estimates of SCALE suggest there are economies of scale in this industry

Table 23.2 Maximum likelihood estimates

Parameter	Estimate (asymptotic "t"-value)	Parameter	Estimate (asymptotic "t"-value)
α_0	7.775 (0.93)	α_1	-0.060 (-0.16)
α_1	0.661 (0.75)	α_2	-1.170 (-2.91)
γ_{yy}	0.031 (0.58)	α_3	2.155 (4.68)
γ_{y1}	-0.004 (-0.51)	β_1	-0.148 (-0.08)
γ_{y2}	-0.018 (-1.59)	β_4	0.267 (0.46)
γ_{y3}	0.023 (2.05)	β_{44}	0.001 (0.02)
γ_{22}	0.197 (8.01)	β_{14}	-0.068 (-0.88)
γ_{32}	-0.168 (-7.17)	β_{17}	-0.037 (-0.24)
γ_{42}	-0.013 (-0.42)	β_{4y}	-0.002 (-0.06)
γ_{33}	0.183 (4.64)	θ_{11}	-0.010 (-0.39)
γ_{43}	-0.000 (-0.02)	θ_{12}	0.150 (3.20)
γ_{41}	-0.018 (-0.83)	θ_{13}	-0.126 (-3.33)
		θ_{41}	0.002 (0.41)
Equation:	In C	S_1	S_2
R^2	0.964	0.661	0.485
		S_3	0.433

Note: Estimates of remaining parameters are derivable from the homogeneity and symmetry restrictions. Asymptotic "t"-values (in parentheses) are standard normally distributed. The asymptotic standard errors on which these are based are calculated using White's heteroskedasticity-consistent covariance matrix estimator.

numerically, all of these values exceed unity, except for that of Auckland, the ESA with by far the largest output. The Wald test is used to test the hypothesis of unitary economies of scale. In fact, we test if the reciprocal of SCALE (from (5)) is unity. The hypothesis of interest is formulated in this equivalent way to minimize the distortion of the true size of the Wald test (from the nominal size chosen) in finite samples. This choice is based on the results of Gregory and Veall (1985) and Phillips and Park (1988). The results obtained appear in table 23.3, where we see that for 26 cases the SCALE estimates are significantly greater than unity. No such estimate is significantly less than unity - even at the level of output experienced by the

Table 23.3 Scale economy measures

Supply authority	Output (10^9 kWh)	SCALE	SCALE
1 Ashburton	0.158	1.090 (3.85)**	1.093 (1.35)
2 Auckland	3.426	0.988 (0.01)	1.016 (0.02)
3 Bay of Islands	0.194	1.082 (3.11)*	1.080 (1.12)
4 Bay of Plenty	0.402	1.057 (0.70)	1.067 (0.43)
5 Bluff	0.017	1.177 (1.41)	1.146 (1.36)
6 Buller	0.082	1.114 (3.43)**	1.117 (1.46)
7 Cambridge	0.087	1.112 (3.56)**	1.119 (1.42)
8 Central Canterbury	0.427	1.055 (0.61)	1.066 (0.50)
9 Central Hawkes Bay	0.082	1.114 (3.44)**	1.101 (1.59)
10 Central Waikato	0.486	1.050 (0.44)	1.061 (1.05)
11 Christchurch	1.475	1.014 (0.01)	1.012 (0.01)
12 Dannevirke	0.070	1.120 (3.08)**	1.124 (2.11)
13 Dunedin	0.731	1.037 (0.15)	1.018 (0.02)
14 Egmont	0.191	1.082 (3.17)**	1.086 (2.03)
15 Franklin	0.248	1.073 (2.06)	1.075 (1.83)
16 Hamilton	0.136	1.095 (4.10)**	1.082 (2.77)*
17 Hawkes Bay	0.565	1.045 (0.30)	1.058 (0.65)
18 Horowhenua	0.257	1.072 (1.91)	1.094 (2.28)
19 Hutt Valley	0.967	1.027 (0.07)	1.058 (0.71)
20 Invercargill	0.216	1.078 (2.65)	1.073 (5.03)**
21 Kaipoi	0.022	1.166 (1.57)	1.162 (1.45)
22 King Country	0.086	1.112 (3.54)**	1.092 (1.20)
23 Manawatu-Oroua	0.306	1.066 (1.32)	1.076 (1.43)
24 Marlborough	0.178	1.085 (3.46)**	1.091 (1.04)
25 Napier	0.076	1.117 (3.26)**	1.122 (2.64)
26 Nelson	0.109	1.103 (4.00)**	1.105 (6.30)**
27 New Plymouth	0.272	1.070 (1.71)	1.050 (0.22)
28 North Auckland	0.417	1.055 (0.64)	1.062 (0.64)
29 North Canterbury	0.198	1.081 (3.03)**	1.082 (0.78)
30 Otago	0.157	1.090 (3.86)**	1.079 (0.78)
31 Otago Central	0.204	1.080 (2.90)**	1.083 (0.61)
32 Palmerston North	0.180	1.085 (3.42)**	1.072 (2.91)*
33 Port Hills	0.079	1.115 (3.36)**	1.106 (3.48)*
34 Poverty Bay	0.207	1.080 (2.84)**	1.106 (3.48)*
35 Riccarton	0.049	1.134 (2.42)	1.147 (1.57)
36 Rotorua	0.285	1.068 (1.55)	1.080 (1.65)
37 South Canterbury	0.260	1.072 (1.87)	1.079 (0.68)
38 Southland	0.424	1.055 (0.62)	1.046 (0.14)
39 Taranaki	0.250	1.073 (2.03)	1.078 (0.98)
40 Taranaki	0.046	1.136 (2.33)	1.134 (1.69)
41 Tasman	0.297	1.067 (1.42)	1.080 (0.76)
42 Taurarunui	0.021	1.169 (1.53)	1.140 (1.94)
43 Taupo	0.162	1.089 (3.78)**	1.058 (0.10)
44 Tauranga MED	0.056	1.128 (2.65)	1.120 (1.55)
45 Tauranga EPB	0.349	1.061 (0.98)	1.075 (1.69)
46 Te Awamutu	0.102	1.106 (3.89)**	1.111 (2.02)
47 Thames Valley	1.062	1.024 (0.05)	1.029 (0.08)
48 Thames-Coromandel	0.030	1.154 (1.82)	1.115 (1.51)
49 Timaru	0.120	1.100 (4.09)**	1.108 (3.43)*
50 Wairarapa	0.195	1.082 (3.08)**	1.089 (1.43)
51 Wairoa MED	0.019	1.171 (1.49)	1.159 (1.42)

continued

Table 23.4 Estimated elasticities

(i, j)	Labor			Capital			Electricity			Other		
	max	min		max	min		max	min		max	min	
Labor	15.76 (14.05)	-4.10* (1.31)										
Capital	0.19 (0.75)	0.64* (0.33)	-0.10 (0.14)	-0.26* (0.14)								
Electricity	-0.43 (3.20)	0.69 (0.70)	0.10 (0.13)	0.06 (0.15)	-0.31 (0.43)	-0.34 (0.22)						
Other	-10.12 (13.48)	-2.54 (4.29)	0.58 (1.00)	0.98 (1.52)	0.97 (1.50)	0.98 (1.12)	-6.46 (17.74)	-6.14 (19.28)				
Cost elasticities of factor demand (η_{ij})												
(i, j)	Labor			Capital			Electricity			Other		
	max	min		max	min		max	min		max	min	
Labor	0.52 (0.46)	-0.44* (0.14)	0.12 (0.46)	0.27* (0.14)			-0.13 (0.97)	0.29 (0.30)		-0.50 (0.66)	-0.12 (0.20)	
Capital	0.01 (0.02)	0.07* (0.04)	-0.06 (0.04)	-0.11* (0.06)			0.03 (0.04)	0.03 (0.05)		0.03 (0.05)	0.02 (0.07)	
Electricity	-0.02 (0.11)	0.07 (0.08)	0.03 (0.05)	0.03 (0.05)	-0.09 (0.13)	-0.15 (0.09)						
Other	-0.33 (0.44)	-0.28 (0.46)	0.36 (0.62)	0.15 (0.64)	0.29 (0.45)	0.41 (0.47)	-0.32 (0.87)	-0.29 (0.91)				

Note: "max" and "min" relate to firms with the largest and smallest outputs in the sample. Figures in parentheses are asymptotic standard errors based on White's heteroskedasticity - consistent covariance matrix estimator.

* Significant at the 10 percent level.

cost relative to that which prevailed in our 1986/87 sample. Even reducing the number of firms from 60 to 40 implies an 8 percent reduction in total industry cost on this basis.⁹

Such figures must be treated cautiously. Our model has been specified carefully but it cannot capture the full detail of this industry. The firms in it are not homogeneous. Each has its individual features, which of course preclude the attainment of the hypothetical situations just described. We offer no prescription for the amalgamation of Supply Authorities. Clearly, there are important geographic, technical, and social constraints that would have to be taken into account. Since this study was undertaken, several major changes to the New Zealand electricity distribution industry have been announced and the final impact of these changes is unclear. Subject to the constraints noted, our results suggest the likely future shape of the industry if Supply Authorities were to rationalize their activities in a more competitive environment.

Notes

The second author is now with Transpower New Zealand Ltd. This paper is based on work undertaken in the preparation of the more general report by Wyatt et al. (1989). We would like to thank Patrick Caragata for suggesting this study and for his ongoing support during

its execution. We are also grateful to Max Brown and Andrew Duncan for their substantial input, to Judith Giles for numerous helpful discussions, and to the New Zealand Ministries of Energy and Commerce for supporting this study. We would like to thank members of the Electricity Supply Association of New Zealand for their many helpful comments and suggestions during the course of our research, and the referee for several constructive comments on an earlier draft. This paper is the responsibility of the authors, and any views expressed should not be attributed to their employers.

1. Six major industrial users purchase bulk electricity direct from the national grid. They are excluded from this study.
2. Recently, this number has been reduced to 52, and negotiations for further amalgamations are in progress.
3. Recent changes to the industry include the potential for ESAs to bargain directly with ECNZ over the price of bulk electricity and changes to taxation arrangements for ESAs. In addition, the New Zealand Government has recently announced that ESAs will be corporatized, so in future they should be more profit-oriented.
4. All of our analysis was repeated with the Auckland observation excluded from the sample. The numerical estimates changed only slightly from those reported below, and none of the conclusions was altered.
5. In this case the implications of the estimation results are, of course, independent of the figure chosen.
6. There is evidence of multicollinearity, in terms of both the simple and multiple correlations between this variable and the other regressors.
7. As the estimate of γ in table 23.2 is positive, the SCALE estimates decrease monotonically with increasing output. This is not the case for SCALE, of course.
8. That paper also reports some tentative, and only partially successful, attempts to model the ESAs as dual-output firms, supplying both "energy" (in kWh) and "power" (in kW). This was motivated by the peak-load problem and the ESAs' limited ability to smooth their loads. The aggregate scale economy estimates obtained were totally consistent with those reported here.
9. In the case of our three-input model these industry cost savings are estimated to be 21 percent and 17.5 percent respectively.

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Long-run Equilibrium Estimation and Inference: A Nonparametric Application

V. B. Hall and R. G. Trevor

Introduction

Phillips (1991a) has demonstrated that the long-run parameters of a continuous-time error correction model (ECM) involving nonstationary variables can be estimated from a corresponding discrete-time ECM. He derives theoretical results for a first-order stochastic differential equation system, driven by quite general stationary errors. The vector of variables is integrated of order one (an $I(1)$ process), with the multiple long-run relationships being given by the cointegrating vectors. This theoretical work is an extension of the Phillips (1991c) results for a discrete-time ECM with a single cointegrating relationship.

He suggests frequency domain procedures of the type due to Hannan (1963) for estimation and inference, anticipating they would provide significant computational advantages over methods traditionally used for continuous-time models.¹ In particular, he considers the very significant problems associated with temporal aggregation (see for example Bergstrom (1984)), the problems of selecting short-run dynamics, and the complexities of nonlinear estimation would not arise. This is due to the generality afforded by the nonparametric treatment of regression errors in frequency domain procedures.

A further key element of Phillips' paper is that spectral estimates of the cointegrating parameters are asymptotically equivalent to the corresponding maximum likelihood estimates. Moreover, the nuisance parameters introduced into the limiting distributions by the presence of nonstationary processes have only scale effects. It is therefore possible to carry out conventional asymptotic chi-squared hypothesis testing.

In light of the above, the principal aim of this paper is to present some initial estimates and inferences based on Phillips' theoretical results and suggested empirical procedures. The innovative aspects are: (1) it provides an early application of this methodology to macroeconomic time-series data,² thereby giving initial evidence on both the potential gains and difficulties relative to traditional methods; (2) it extends the Engle and Granger (1987) type scalar cointegration methods to vector cases, in a way which is not conditional on the precise modeling of short-run dynamics as is required, for example, in the maximum likelihood procedure developed by Johansen (1988, 1991); and (3) it illustrates the outcome of some simple hypothesis tests on

TABLE 7 (PART 2): - ANALYSIS OF CONSUMPTION OF ELECTRICITY PER SECTOR FOR THE YEAR ENDED MARCH 1990

	INDUSTRIAL				NON-DOMESTIC				TOTAL RETAIL SALES							
	Electricity Consumed kwh	Income \$	Income per kwh sold c	No. of Consumers Number	Average kWh per consumer	Electricity Consumed kwh	Income \$	Income per kwh sold c	No. of Consumers Number	Electricity Consumed kwh	Income \$	Income per kwh sold c	No. of Consumers Number			
1 Electropac	6658061022	238817982	3.59	5	1331612204	6704605311	246218295	3.67	6	494219	6,705,372,150	246287213	3.67	83	787616	1
13 Waioia Electricity	568554	80078	14.13	4	141639	7138838	1066413	14.94	351	20330	19,620,735	2415008	12.31	2142	9160	13
4 Kaipoi Electricity	4149994	362035	8.72	27	153703	8157298	819208	10.04	203	40784	23,688,400	1952600	8.29	3466	12049	4
32 Waioia EPB	20175505	2010781	9.97	16	1260969	24586465	2815238	11.45	1174	20942	38,439,825	4249227	11.05	3466	11094	32
8 Tauranga DECD	8034676	791960	9.78	121	66898	536316	536316	11.74	1667	27402	61,436,102	6867012	11.18	4247	14466	11
8 Dannevirke EPB	1443857	1439082	9.97	36	401071	30286649	3646306	12.04	2365	12806	67,032,109	6918569	10.31	6971	9624	8
6 Napier City Council	7281559	932083	12.63	22	335525	45028604	5903330	13.24	1195	37679	75,665,280	8891822	11.75	5670	13345	6
7 Central Hawkes Bay EPB	19014873	1643948	8.65	43	442206	34285552	3702244	10.8	2018	16990	77,809,992	7656871	9.84	7286	10679	7
5 Buller Electricity	50536356	2740076	5.42	55	918843	59893857	3915015	6.56	800	74617	78,767,618	5586863	7.1	4231	18617	5
14 Whangarei DECD	6759676	828052	12.25	91	74282	50104221	6672334	13.32	1617	30986	92,360,818	10267500	11.12	8209	11251	14
9 Port Hills Energy	30195919	1907274	6.32	81	3774490	60397662	4235393	7.01	367	164571	100,115,501	7000364	6.99	10041	23991	9
28 Te Awamutu EPB	4850910	743287	15.32	36	134748	41532539	5118911	12.33	3144	13210	102,219,675	10468107	10.24	10269	9954	28
6 Cambridge EPB	35651541	2582328	7.24	32	1114111	56352103	5334493	9.47	1964	28693	102,697,172	9704710	9.45	7632	13456	6
1 Clipponer Nelson	19470867	1639742	8.42	65	289553	59914493	6487236	10.78	1637	36600	109,114,807	10238244	9.38	1800	13146	1
15 King Country EPB	10707557	873261	8.16	53	202029	51627369	5195979	9.99	3134	16473	111,739,994	8982212	8.02	1803	9465	15
35 Waitomo EPB	30766367	3261873	10.6	81	379832	59798526	6883967	11.51	4259	14041	116,253,112	12680776	10.91	11989	9713	35
12 Timaru Electricity	13939457	1338977	8.17	50	327869	59547705	5421137	9.1	1482	40181	121,110,609	9985147	8.24	8757	13830	12
37 Westport	27990336	2950725	10.54	28	207683	78403354	8610365	10.98	2592	30248	132,748,843	13771292	10.37	11394	11651	37
20 Otago EPB	31955945	2564696	8.03	76	409692	64461481	7039719	10.92	3979	16200	146,467,453	14538718	9.93	14624	12707	20
33 Waitaki EPB	82821735	6492431	7.84	52	1592726	121551284	11719422	9.64	3173	38308	185,403,334	17846665	9.63	12049	10279	10
10 Egmont EPB	9849730	1332223	13.53	99	99492	69007067	12146479	13.65	2818	31585	187,690,961	19105684	10.18	18256	10378	8
8 Palmerston North MED	93764378	6511858	6.94	43	2181032	132697630	11757019	8.72	1786	73885	195,359,257	17085279	8.74	10637	18366	10
31 Wairarapa EPB	19300374	2857181	13.25	194	99486	76881573	9869261	12.84	4246	18107	196,988,469	22105659	11.22	19124	10301	31
1 Ashburton	39597858	2989340	7.55	27	1466587	113157567	9276244	8.2	4708	30517	198,237,084	16588063	8.35	12107	16374	1
17 Maniborough Electric	33324944	3105267	9.32	174	191523	92000739	9246738	10.55	3305	27804	201,098,052	17694897	8.8	17903	11233	17
9 Eastland Energy	44154285	5119669	11.59	85	519462	92802463	13492369	14.54	5126	18104	202,441,495	25139459	12.42	21229	9536	9
3 Bay of Islands EPB	37319189	3438919	9.21	54	691096	97696070	12106751	12.39	4805	21687	208,190,872	22545492	10.83	20964	9931	3
21 Otago Central EPB	16639277	1142533	6.87	30	554643	107652077	9334211	8.67	3109	34628	214,830,734	16847644	7.84	19746	13644	21
19 North Canterbury EPB	58739950	3797779	6.47	102	575882	113573219	9379860	8.26	4412	25742	225,715,588	18486131	8.19	18552	12167	19
3 Invercargill MED	20226132	1820509	9.12	129	156792	103125334	10439563	10.12	2656	38827	240,110,394	20949217	8.72	17628	13621	3
24 South Canterbury EPB	64479002	4563690	7.08	43	1499512	118792352	9110828	7.67	3750	31678	251,934,463	18412445	7.31	19157	13151	24
25 Teranaki EPB	137140607	10248831	7.47	125	1097125	174678668	15206673	8.71	3824	45680	284,996,887	23091463	9.06	14516	17567	25
11 Franklin EPB	56578534	5763035	10.19	260	217610	128086199	15211259	11.88	7715	11602	272,888,298	28508040	10.45	26634	10246	11
13 Horowhenua EPB	11938372	1661368	13.92	84	142123	96510628	14029589	14.64	5082	18991	276,736,472	30776381	11.12	34818	7948	13
7 New Plymouth Energy	66209095	3603753	5.44	286	231500	148650793	14857553	10.2	4135	35224	277,653,171	27597293	9.94	24044	11548	7
22 Rotorua Electricity	31945916	3703155	11.59	117	273042	138367999	17564347	12.69	4881	28495	307,885,316	32696989	10.62	27392	11240	22
26 Tasman Energy	99226898	8315683	8.38	503	197270	174089690	17291714	9.93	4793	36322	314,492,329	28608289	9.1	24228	12981	26
36 Wanganui-Hangitikei EPB	48881170	5460345	11.17	116	421389	150549382	17661682	11.73	7120	21145	334,419,282	33394684	9.99	33184	10078	36
16 Manawatu-Taranui EPB	108740666	10874788	9.63	199	286970	152288879	16461414	10.81	7693	19796	381,840,093	36885728	10.13	43158	8847	27
27 Tauranga EPB	56867717	5483378	9.27	137	893723	208616192	21793992	10.94	10375	19915	411,630,859	39166926	9.52	34883	11767	1
4 Bay of Plenty EPB	298640246	18761155	6.28	114	4666254	364765533	26622608	7.3	5006	72864	489,696,351	37105695	7.7	22029	22230	4
18 North Auckland EPB	305516267	19740329	6.46	62	272724	377737613	28564120	7.56	6220	60730	534,318,804	42721422	8	30175	17707	18
12 Hawkes Bay EPB	139529002	13120486	9.4	213	655066	267461707	29379215	10.98	7139	28859	648,653,610	68070177	10.49	45498	12364	12
30 Waikato Electricity	106262366	9577315	9.01	437	343163	326311511	37330650	11.44	11307	28859	648,653,610	68070177	10.49	45498	12364	12
2 Dunedin Electricity	88926440	6826432	7.68	255	348770	301946212	24729121	8.19	6645	45440	737,642,010	51524332	6.99	49005	15052	2
14 Hutt Capitalpower	33439051	3887739	10.73	190	175995	525053689	56029444	11.95	8367	67584	918,119,155	82619031	9.11	66915	16131	5
5 Mtit Valley Energy Board	160621712	16371115	10.19	448	358531	427505426	51166388	11.97	8332	61294	990,847,693	96646283	9.75	82688	11983	14
29 Thames Valley EPB	716102807	46890244	6.52	283	2530399	867249270	66718889	7.73	11552	74684	1,131,005,589	91143396	8.06	54467	20765	29
34 Waitemata EPB	148457644	15115737	10.18	959	154805	498085793	60585856	12.13	17000	29358	1,433,860,147	138710988	9.67	143869	9966	34
23 Southpower	277797543	25424918	9.15	1019	272618	895940345	90821673	10.13	17227	52066	2,013,768,831	169188657	8.4	132041	15251	23
2 Auckland EPB	618789332	49778762	8.04	925	668973	2197818501	227806731	10.37	31584	69586	3,712,879,052	346225504	9.32	230314	16121	2

TABLE 7 (PART 2): - ANALYSIS OF CONSUMPTION OF ELECTRICITY PER SECTOR FOR THE YEAR ENDED MARCH 1990

	INDUSTRIAL				NON-DOMESTIC				TOTAL RETAIL SALES			
	Electricity Consumed kwh	Income \$	No. of Consumers	Average kWh per consumer	Electricity Consumed kwh	Income \$	No. of Consumers	average kW per consumer	Income \$	Electricity Consumed kwh	No. of Consumers	average kWh per consumer
1 Electricorp	6658061022	238817982	3.59	1331612204	6704605311	248218295	6	434219	248287213	3.67	83	787616
1 Ashburton	39597858	2989340	7.55	1466587	113157567	9276244	3708	30517	198237084	8.35	12107	16374
2 Auckland EPB	618799732	49778762	8.04	688973	2197818601	227806731	31584	69586	3712879052	9.32	230314	16121
3 Bay of Islands EPB	37319189	3438919	9.21	691096	97699070	2106751	4505	21687	208190872	10.83	20964	9931
4 Bay of Plenty EPB	68640246	8761155	6.28	4466254	364755533	266226018	7.3	5006	489696351	7.7	22029	22200
5 Buller Electricity	50836356	2740076	5.42	918843	59639515	3915015	6.56	800	74617	78767618	4.23	18617
6 Cambridge EPB	35851541	2682328	7.24	1114111	563521103	5334293	9.47	1964	28693	102697172	9.45	7632
7 Central Hawkes Bay EPB	19014873	1643948	8.65	442206	34285532	3702244	10.8	2018	16990	77809392	9.84	7286
1 Central Nelson	19470967	1639742	8.42	299553	59914993	6467236	10.79	1637	36600	109114807	9.38	8900
8 Dannevirg EPB	14438557	1439082	9.97	401071	30286649	3648306	12.04	2365	12806	67092109	10.31	6971
2 Dunedin Electricity	88936440	6828542	7.68	348770	301946212	24729121	8.19	6645	45440	737642010	6.99	49005
9 Eastland Energy	44154285	5119659	11.59	1592726	92802463	13492369	14.54	5126	18104	202441495	12.42	21229
10 Egmont EPB	82821735	6492431	7.84	52	121581284	11719422	9.64	3173	38308	185403334	9.63	12049
11 Franklin EPB	56578534	5763035	10.19	217610	128086199	18211259	11.88	7715	16602	272888298	10.45	26634
12 Hawkes Bay EPB	139529002	13120486	9.4	655066	267461707	29379215	10.98	7139	37485	562540413	9.74	45498
13 Horowhenua EPB	11938372	1661368	13.92	142123	96510528	14029959	14.54	5082	18991	276736472	11.12	34818
14 Hutt Valley Energy Board	20226132	16371115	10.19	448	103125334	10439563	10.12	2656	38827	240110394	8.72	17628
3 Invercargill EPB	66209095	362035	8.72	153703	8152798	8192078	10.04	203	40184	233688400	8.29	1966
4 Kaipoi Electricity	1070557	873261	8.16	53	51627369	515979	9.99	3134	16473	111739994	8.02	11805
15 King Country EPB	108740666	10874788	10	293	202029	21739818	10.96	7154	2739	364627100	9.6	30047
16 Manawatu-Taranaki EPB	33249444	3105267	9.32	171	92000739	99602944	11.35	8367	62754	918119155	9.11	56915
5 MED Capitalpower	33439051	3587739	10.73	190	525063689	59602944	13.24	1195	37679	75665280	8891822	13345
6 Napier City Council	7381559	932083	12.63	335525	45026804	14857553	10.2	4135	35224	2776653171	9.94	24044
7 New Plymouth Energy	66209095	362035	5.44	286	145606793	14857553	10.2	4135	35224	2776653171	9.94	24044
18 North Auckland EPB	305516267	19740329	6.46	112	37737813	28564128	7.56	6220	60730	534318804	4.27	30175
19 North Canterbury EPB	58739560	3797779	6.47	102	113573219	9379860	8.26	4412	25742	225715588	8.19	18552
20 Otago EPB	16639277	1142533	6.87	30	107652077	9334211	8.67	3109	34626	214830734	7.84	15746
9 Palmerston North MED	27990336	2950725	10.54	26	64461481	7039719	10.92	3979	16200	146467453	9.93	14706
8 Port Hills Energy	30195919	1907274	6.32	8	89007067	12146479	13.65	2818	15885	187660981	10.18	18256
22 Rotorua Electricity	64479002	4563690	7.08	43	118792352	9110828	12.69	4861	28465	307885316	10.62	27392
23 Southland EPB	122440100	11345087	9.15	1019	206616192	21778921	10.54	10375	19915	411630859	9.52	34983
25 Taranaki EPB	99226698	8315683	8.38	503	896940346	90821673	10.13	17227	52066	2013788831	8.4	13204
10 Taupo Electricity	93784378	6511858	6.94	43	174089690	17291714	9.93	4793	36322	314492329	9.06	14516
11 Tauranga DCED	6969717	5483378	9.78	121	132667830	11575019	8.72	1796	73885	195359257	8.74	10637
28 Te Awamutu EPB	4850910	743287	15.32	36	45678840	5363616	11.74	1667	27402	614365202	11.18	4247
29 Thames Valley EPB	716102807	4690244	6.52	283	152288879	16461414	10.81	7693	19796	381840003	38685728	1013
30 Waikato Electricity	16393457	1338977	8.17	50	862749270	6718888	12.33	3144	13210	1022191675	10.24	10269
31 Waipara EPB	19300374	2557181	13.25	194	326311511	37330650	11.44	11307	28859	648653610	68070177	1049
32 Waipa EPB	20175805	2010781	8.97	16	76881573	9866213	12.84	4246	18107	196988489	11.22	19124
33 Waitaki EPB	13995845	2564696	8.03	76	11353538	1066413	14.94	351	20300	19620735	2415008	1231
34 Waitmatia EPB	148457644	1511737	10.18	959	24586465	2815238	11.45	1174	20942	38439825	38439825	31
35 Waitemata EPB	30766367	3261873	10.6	81	499085793	6059589	12.13	17000	29322	147706631	11855668	8.03
36 Wanganui-Rangitikei EPB	48881170	5460345	11.17	116	150549326	17661682	11.73	7120	21445	334419262	12680776	10.91
37 Westpower	35513804	3659838	10.31	171	78403344	8610365	10.98	2592	30248	132748843	13771292	10.37
14 Whangarei DCED	6759676	828052	12.25	91	50104221	6672334	13.32	1617	30986	92360818	10267500	11.12